

Foglio Fronte/Retro 10x15 per esame di Analisi Reale.

$X: S \rightarrow R$	R	$P(X=n) f_X(t)$	$E(X)$	$Var(X)$
$B(p)$	0,1	p	p	$1-p$
Binomiale $B(n, p)$	$\{0, 1, \dots, n, \dots\}$	$\binom{n}{k} p^k (1-p)^{n-k}$	np	$np(1-p)$
Poiss(λ)	$\{0, 1, \dots, n, \dots\}$	$e^{-\lambda} \lambda^n / n!$	λ	λ
$N(\mu, \sigma^2)$		$\frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2}(\frac{t-\mu}{\sigma})^2}$	μ	σ^2
exp(λ)		$\lambda e^{-\lambda} \quad t \geq 0$ $0 \quad t < 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
$\sigma(p)$	$\{1, 2, \dots, n, \dots\}$	$p(1-p)^{n-1}$		

$$P(A) = P(A \cap B_1) + \dots + P(A \cap B_n) = P(A|B_1)P(B_1) + \dots + P(A|B_n)P(B_n)$$

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \quad P(B_k|A) = \frac{P(A|B_k)P(B_k)}{P(A|B_1)P(B_1) + \dots + P(A|B_n)P(B_n)}$$

$$E(X) = \int_{-\infty}^{+\infty} t f_X(t) dt \quad E(X) = \sum_{n=1}^{\infty} x_n p_n \quad S^2 = \frac{SS}{n-1} = \frac{1}{n-1} \sum_{i=1}^n (X_i - M_n)^2$$

$$Var(X) = E((X - E(X))^2) = E(X^2) - E^2(X) \quad E(SS) = (n-1)\sigma^2$$

$$Var(X+Y) = Var(X) + Var(Y) + 2[E(XY) - E(X)E(Y)] \quad M_n = \frac{X_1 + \dots + X_n}{n} \quad E(M_n) = \mu$$

$$\lim_{n \rightarrow \infty} P(a \leq X_n \leq b) = \int_a^b \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt = \Phi(b) - \Phi(a) \quad Var(M_n) = \frac{\sigma^2}{n}$$

$$S_{XY} = \frac{1}{n-1} \sum_{i=1}^n (X_i - M_X)(X_i - M_Y) \quad \rho_{XY} = \frac{COV(X, Y)}{\sigma_X \sigma_Y} \quad a = M_X - b M_Y$$

$$COV(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] = E(XY) - E(X)E(Y) \quad Z = \frac{X - \mu}{\sigma} \quad b = \frac{S_{XY}}{S_X^2}$$

Q.Pivotale (QP)	Metodo	I_α
$S_n = \frac{\sqrt{n}}{\sigma} (M_n - \mu)$ Lim. Centrale	La QP tende in prob alla Normale STD	$P(-z_{\alpha/2} < S_n < z_{\alpha/2}) = 1 - \alpha$ $I_\alpha = \left(M_n - \frac{\sigma}{\sqrt{n}} z_{\alpha/2}; M_n + \frac{\sigma}{\sqrt{n}} z_{\alpha/2} \right)$
$S_n = \frac{\sqrt{n}}{S} (M_n - \mu)$	La QP è la T_{n-1} di Student	$P(-q_{\alpha/2} < S_n < q_{\alpha/2}) = 1 - \alpha$ $I_\alpha = \left(M_n - \frac{S}{\sqrt{n}} q_{\alpha/2}; M_n + \frac{S}{\sqrt{n}} q_{\alpha/2} \right)$
$C_{n-1} = \frac{n-1}{\sigma^2} S^2$ Teo. Fisher	La QP si comporta per $n \rightarrow \infty$ come la χ_{n-1}^2	$P(q_{\alpha/2} < C_{n-1} < q_{1-\alpha/2}) = 1 - \alpha$ $I_\alpha = \left(\frac{(n-1)s^2}{q_{1-\alpha/2}}; \frac{(n-1)s^2}{q_{\alpha/2}} \right)$

Per Paerson sappiamo che la D_{N-1} tende alla χ_{N-1}^2 per n grande.

N Num. valori assunti dalla Y n Num. campioni della X
 n_i Freq. della Y in base alla sua legge e riferite a $n \Rightarrow n_i = n p_i$
 v_i Freq. che il campione assume val. $y_i \Rightarrow v_i = Card\{x_j | x_j = y_i\}$

$$D_{N-1} = \sum_{i=1}^N \frac{n}{p_i} \left(\frac{v_i}{n} - p_i \right)^2 = \sum_{i=1}^N \frac{(v_i - n p_i)^2}{n p_i} \quad I_\alpha = (0, q_{1-\alpha})$$

Cebicev $P\{|X - \mu| \geq \eta\} \leq \frac{\sigma^2}{\eta^2} \int_{-\infty}^{\infty} (t - \mu)^2 f_X(t) dt$

Numeri grandi $P\{|M_n - \mu| \geq \eta\} \leq \frac{\sigma^2}{n \eta^2}$

$$E(SS) = E \left[\sum_{i=1}^n (X_i - \mu + \mu - M_n)^2 \right] = \sum_j \sum_k (x_j + y_k) P_\xi(x_j, y_k) = \sum_j \sum_k (x_j + y_k) P_\xi(x_j, y_k) = \sum_j \left(\sum_k x_j P_\xi(x_j, y_k) \right) + \sum_k \left(\sum_j y_k P_\xi(x_j, y_k) \right) = \sum_j x_j P_{x_j} + \sum_k y_k P_{y_k}$$

$$P\{X \in I_1, Y \in I_2\} = P\{X \in I_1\} P\{Y \in I_2\} \Rightarrow \int_{I_1, I_2} f_\xi(t, s) dt ds = \int_{I_1} f_X(t) dt \int_{I_2} f_Y(s) ds$$

$$\hat{\theta} \rightarrow \theta \text{ in prob.} \Leftrightarrow \lim_{n \rightarrow \infty} P\{|\hat{\theta} - \theta| > \eta\} = 0$$

$$A = \{s \in S \text{ t.c. } |\hat{\theta} - \theta| \leq \eta\} \quad B = \{s \in S \text{ t.c. } |\hat{\theta} - E(\hat{\theta})| \leq \eta/2\}$$

se $n > n_0 \Rightarrow B \subset A \quad \underbrace{|\hat{\theta} - \theta|}_{\eta} \leq \underbrace{|\hat{\theta} - E(\hat{\theta})|}_{\eta/2} + \underbrace{E(\hat{\theta}) - \theta}_{\eta/2}$

Cebicev $\Rightarrow P\{|\hat{\theta} - E(\hat{\theta})| > \eta/2\} \leq \frac{Var(\hat{\theta})}{(\eta/2)^2} \rightarrow 0 \text{ se } n \rightarrow \infty$

$\forall \epsilon > 0 \exists n_1 \text{ t.c. se } n \geq n_1 \Rightarrow P\{s \in S \text{ t.c. } |\hat{\theta} - E(\hat{\theta})| > \eta/2\} < \epsilon$

$P\{s \in S \text{ t.c. } |\hat{\theta} - \theta| \leq \eta\} \geq 1 - \epsilon$

se $n > \max(n_0, n_1) \Rightarrow P(B) \geq 1 - \epsilon \Rightarrow P(A) \geq 1 - \epsilon$

$\Rightarrow \forall \epsilon > 0 \exists n_2 = \max(n_0, n_1) \text{ t.c. se } n \geq n_2 \Rightarrow P(A) \geq 1 - \epsilon$

$\lim_{n \rightarrow \infty} P\{|\hat{\theta} - \theta| \leq \eta\} = 1 \Rightarrow \lim_{n \rightarrow \infty} P\{|\hat{\theta} - \theta| > \eta\} = 0$

$$f_{T_n}(t) = \frac{\Gamma((n+1)/2)}{\Gamma(n/2) \sqrt{n\pi}} \left(1 + \frac{t^2}{n}\right)^{-(n+1)/2} \quad f_{\chi_n^2} = \begin{cases} \frac{(1/2)^{n/2}}{\Gamma(n/2)} e^{-t/2} t^{n/2-1} & t > 0 \\ 0 & t \leq 0 \end{cases}$$

$$r_{XY} = \frac{\frac{1}{n-1} \sum_{i=1}^n (X_i - M_X)(Y_i - M_Y)}{S_X S_Y} \quad Z = \frac{\sqrt{n-3}}{2} \ln \left[\frac{(1+r_{XY})(1-\rho_{XY})}{(1-r_{XY})(1+\rho_{XY})} \right]$$